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A Modified Approach by Using Prediction to Build a Best Threshold in ARX Model with Practical Application

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Abstract:

The proposal of nonlinear models is one of the most important methods in time series analysis, which has a wide potential for predicting various phenomena, including physical, engineering and economic, by studying the characteristics of random disturbances in order to arrive at accurate predictions.

In this, the autoregressive model with exogenous variable was built using a threshold as the first method, using two proposed approaches that were used to determine the best cutting point of [the predictability forward (forecasting) and the predictability in the time series (prediction), through the threshold point indicator]. B-J seasonal models are used as a second method based on the principle of the two proposed approaches in determining the best seasonal model. Then they are compared with the obtained models from two methods that mentioned above of the two approaches within a group of the criteria as AIC, MDL, Loss Function, BIC, FPE, MSE, in addition the proposed weighted comparison criteria to determine the best model for representing the wind speed data as input variable, soil and dust as an output variable, in Baghdad Station from January 1956 to December 2012.

Key words: ARX, Forecasting, Prediction, Seasonal autoregressive integrated moving average, Threshold.

Introduction:

The approaches of analyzing time series are among the most important methods of building models and predicting the various applicable phenomena. Most of the applied data are regarded as nonlinear and are characterized by randomness, and most of the prediction methods may not pay attention to this aspect in conducting and analyzing data. This affects the accuracy of the results which can be obtained. Therefore, all the data possibilities of the various studied phenomena, their quality, their extent of being affected and other factors that may relate to this data should be taken into account, and such considerations should be taken so that the researcher can choose the appropriate model to the nature of this data.

Nonlinear models have been important stages in time series analysis since the end of the 19th century, as they are regarded nonlinear extensions of ARIMA models which contributed effectively to improving predictions of various daily life phenomena by studying the features of nonlinear random disorders which contribute to reaching precise future predictions (1).

The objective of this paper is to build the best model of soils and dust data by the explanatory variable "wind speed" in Baghdad Station. This is done by building the ARX model by using the threshold as a first method, and adopting two proposed approaches employed to determine the best threshold point (Forecasting and Prediction, by the threshold point indicator). The ordinary seasonal Box-Jenkins models were used as a second method, depending on the principle of the two proposed approaches in determining the best seasonal model of data. Further, there will be comparison with the models obtained from the two approaches mentioned above and both methods by a set of statistical criteria to determine the best model in data representation.

The Theoretical Part: Nonlinear Time Series

Kaldor in 1940 and Goodwin in 1955 caused a paradigm shift in the analysis of economic time series (2, 3) by using nonlinear models which depend on time in the analysis of financial and monetary problems because they are characterized by dynamism and movement. In 1958, Weiner continued this way and studied nonlinear variants of ARMA models. Moreover, the classical economists developed the nonlinear ARMA models to identify

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and estimate the nonlinear symmetric and multinomial phenomena as they are duplication models and models of regimes change.

Consequently, many nonlinear time series models were used. Tong (4) proposed in 1978 and 1990 the autoregressive threshold model by precisely describing the periodic symmetrical limit of many annual sunspots. Also, Haggan and Ozeki in 1981 studied the exponential autoregressive model and showed that it is possible to use it in modeling sound vibration. Thus, researchers began analyzing nonlinear time series in exploring the possibility of using several approaches such as non-parametric intensity in modeling nonlinear conduct in the economic, financial and environmental time series, and the like.

In 2010, Lucheroni (5) presented two models for building TARX for the prices of electricity measured in hours for one week, taken from the Canadian Energy Market AESO at the interrupted time and the continuous time. The first model depended on the McKeen model of nerve cells with increasing oscillation. The second model is a generalization of the first by using a method of explaining and showing the random oscillations and fluctuations, taking into account the time when the peak of heights (prices) occurs during daytime, and also explaining their conduct by modeling a mathematical threshold relating to power grid congestions.

In 2012, Yousfat (6) studied the relationship between inflation and economic growth in Algeria during the period (1970-2009) by using the Khan and Senhadji model for 2000 to determine the threshold level of inflation. The study concluded that the threshold level of inflation in Algeria is 6%, i.e., the inflation rates larger than the threshold may harm the economic growth in the country.

In the same year, Filipovic, Stojanovic, Nedic and Prsic (7) presented an engineering study on pneumatic cylinders by putting forward workrelated hypotheses, i.e. the possibility of bringing closer the nonlinear model of the cylinder to timechanging ARX model which is called TARX. Due to the influence of the mixture of the thermal coefficient, coefficient, the discharge temperature, it was supposed that the cylinder parameters are random and that their observations are Gaussian distribution. The random model with changing parameters was supposed, and the algorithm of Kalman filter's approach was used.

The Threshold Model

The threshold models have been developed by economists specializing in econometrics and are called smoothing thresholds developer, which Tong dealt with in 1978. It included several models (8, 9): self-exciting threshold autoregressive (SETAR), self-exciting threshold autoregressive/moving average (SETARMA), smooth threshold autoregressive (STAR), exponential autoregressive (EAR) and open-loop threshold autoregressive system (TARSC), and other models which have benefits in the economic, financial, engineering and environmental fields, and others.

The threshold model is regarded a nonlinear model in time series which have nonlinear approaches in their conduct in a specific period. The point can be determined in many approaches, the most important of which is the separation point which is called the threshold point. This point is based on the values preceding the exogenous variable which determines the shape of the model before and after the threshold point according to the following formula:

$$y_{t} = \begin{cases} b_{1}y_{t-1} + \varepsilon_{t} & \text{if } y_{t-1} < r \\ b_{2}y_{t-1} + \varepsilon_{t} & \text{if } y_{t-1} \ge r \end{cases} \dots (1)$$

Whereas:

 ε_t : Strict white noise.

The above model is one of the SETAR models which can be generalized as follows:

$$y_t = b_0^{(i)} + \sum_{i=1}^{p_i} b_j^{(i)} y_{t-j} + \varepsilon_t$$
 ...(2)

Whereas:

p : Order for (AR) model.

 b_j , b_0 : The parameters for the model, and $i = 1, 2, \dots, k$ represents the number of sections.

 y_{t-j} : Observation values y at point (t-j)

The model in equation (2) is written for short as follows: $SETAR(k, p_1, p_2, \dots, p_k)$

ARX Model

It is an extension of the autoregressive model (AR) with exogenous variables (X) and can be represented in the following formula: (10, 11)

$$y_t = \frac{B(q)}{A(q)}u_t + \frac{1}{A(q)}\varepsilon_t \qquad \qquad \dots (3)$$

Whereas:

 y_t : Outputs at a time (t), and $\frac{1}{A(q)}$ is a linear filter.

A(q): Multinomial, represented as follows: $A(q) = 1 + a_1 q^{-1} + a_2 q^{-2} + \dots + a_{nq} q^{-nq}$

B(q): Multinomial, represented as follows: $B(q) = 1 + b_1 q^{-1} + b_2 q^{-2} + \dots + b_{nb} q^{-nb}$

na, nb: Represent the parameters of multinomials. q^{-i} : represents back shift, and i=1, 2, ...

 u_{\cdot} : represents inputs at time t.

The above equation can be written as follows:

$$A(q)y_t = B(q)u_t + \varepsilon_t \qquad \dots (4)$$

The ARX Model is regarded as one of the transfer function models and is characterized by its applicability to most phenomena and distinguished results. It is also characterized by preparing the two input and output series. The great interest in the use of the ARX model has started particularly in recent years. It is noteworthy that the Mann and Wald studied this model in 1943, whereas in 1960 Durbin studied it. The model contributes to the spread of special functions to obtain high accuracy in the investment management of the results of economic, financial and engineering time series, and others.

Studies in this field continued. In 2000. Knotters and Bierkens (12) dealt with the ARX model by studying the relationship between excessive rainfall and the depth of the underground water level. They studied soil-water balance and the influence of hydrologic intervention of underground water in two sites in the eastern side of the Netherlands and they study the predictions in it as well. The study included two periods, the first in 1985 and the second in (1995-1996).

The researcher used Al-Talib (13) in 2012 as ARX model on data relating to temperature which is affected by a set of explanatory variables, i.e. solar radiation, solar luminosity and evaporation to determine the class of the model by using many criteria, including AIC criterion, FPE criterion and others. Many methods are used to estimate the model parameters, including LS method, IV method and others. Also, the efficiency of the identified model was verified and its predictions were studied.

The Threshold ARX Model

The SETAR model, proved by Kapetanios (14), can be generalized into the threshold ARX model, i.e. SETARX. This can be illustrated in general through the algorithm in Fig.1:

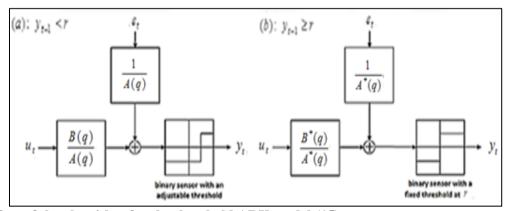


Figure 1. One of the algorithm for the threshold ARX model (15)

Whereas: A(q), $A^*(q)$ are a polynomial of parameters and back shift operator for the variable

B(q), $B^*(q)$ are a polynomial of parameters and back shift operator for the variable u_t .

 ε_t is a prediction error.

Where a simple initial model can be build, as shown

in the following equation:(16)
$$y_{t} = \begin{cases} \theta_{1} \varphi_{t-1} + \varepsilon_{t} & \text{if } y_{t-1} < r \\ \theta_{2} \varphi_{t-1} + \varepsilon_{t} & \text{if } y_{t-1} \ge r \end{cases} \dots (5)$$

Whereas: θ_1 the vector of parameters of the model when $y_{t-1} < r$.

 θ_2 the vector of parameters of the model when $y_{t-1} \ge r$.

represents the inputs and outputs of the φ_{t-1} system:

$$\varphi'_{t-1} = \begin{bmatrix} y_t & y_{t-1} & y_{t-2} & \dots & y_{t-na} & u_t & u_{t-1} & u_{t-2} & \dots & u_{t-nb} \end{bmatrix}'$$

The simple model in equation (5) can be generalized into two groups, each group includes the size of a different sample n_1, n_2 so that there will be more than one separation point (r_1, r_2, \dots, r_k) , where k represents the number of sections separating the two models at a certain point belonging to the points of the depended (exogenous) variable y_{t-l} , and l represents the amount of back shift consequently, the general form of the model will be as follows:

$$y_{t-1} = \theta_0^{(i)} + \sum_{j=1}^{n_{f_i}} \theta_j \varphi_{t-j} + \varepsilon_t$$
 ... (6)

Whereas:

$$n_{f_i} = n_{a_i} + n_{b_i}$$

 n_a . Represents the order of the dependent variable y , and n_{b_i} represents the order of the explanatory variable X, and $i = 1, 2, \dots, k$.

The model in the equation (6) can be summarized as follows: $SETARX(k, n_{f_1}, n_{f_2}, \dots, n_{f_k})$

Comparative Criteria

The criteria below were used to determine the best cut point (threshold) for ARX model by the smallest value of these criteria equivalent to the model. These criteria can also be used to compare between many significant models in the seasonal Bto determine the best significant J approach model.(17, 18, 19)

Akaike's Information Criteria:

Akaike (1973, 1974) used Akaike's information criteria, which is denoted as (AIC), in choosing the suitable order for the models of time series. Its formula is as follows:

$$AIC(M) = -2 \ln L + 2M \qquad \dots (7)$$

Whereas M represents the number of the model's parameters, and L is the likelihood function used to estimate the model.

Bayesian Information Criterion:

The general formula for this criterion, which is denoted (BIC), is:

$$BIC(M) = n \operatorname{Ln} \hat{\sigma}_{a}^{2} - (n - M) \operatorname{Ln} (1 - \frac{M}{n}) + \dots (8)$$

$$M \operatorname{Ln} n + M \operatorname{Ln} \left\{ \left[\frac{\hat{\sigma}_{Z}^{2}}{\hat{\sigma}_{a}^{2}} - 1 \right] / M \right\}$$

Whereas $\hat{\sigma}_Z^2$ represents the estimation of the variance time series, and $\hat{\sigma}_a^2$ represents the variance of the residuals time series, and M represents the number of parameters. Akaike suggested that some terms could be neglected to avoid the overestimate order, then the criterion become:

$$BIC(M) = n \ln \hat{\sigma}_a^2 + M \ln (n) \qquad \dots (9)$$

Final Prediction Error:

This criterion was used in 1969 by the researcher Akaike and denoted as (FPE). It is calculated according to the variance of prediction error for the next period, its formula is:

FPE =
$$[(1 + M / n)/(1 - M / n)] * V$$
 ...(10)
Whereas M represents the number of parameters and V represents the loss function and its formula is as follows:

$$V = \frac{1}{n} \sum e_t^2 \qquad \dots (11)$$

 $V = \frac{1}{n} \sum e_t^2$...(11) Whereas: e_t is prediction error and equal: $e_t = y_t - \frac{1}{n} \sum_{t=0}^{n} e_t^2$

Mean Square Error:

The mathematical formula of this criterion, which is denoted (MSE), equal to:

$$MSE = \frac{\sum (y_t - \hat{y}_t)^2}{n} \qquad \dots (12)$$

Minimize Description Length Criteria:

The mathematical formula of minimizing description length criteria, which is denoted (MDL),

$$MDL = V\left(\frac{1+M \ Log(n)}{n}\right) = V_{\text{mod}} \qquad \dots (13)$$

Whereas V represents the loss function and Mrepresents the number of parameters inside the model, and n represents the size of the sample.

Weighted Comparison Criteria:

In this research, the weighted comparison criteria were proposed using threshold models to determine the general comparison criterion of the data. There is a turning point (cut) where the data will divide into two parts, consequently, a model will be built for each part, and accordingly each part will have a comparison criterion.

If we assume that we have two models: the first has a comparison criterion of, for example, Δ_1 with the size n_1 , while the second has a comparison criterion of, for example, Δ_2 with the size n_2 Therefore, it was proposed to use the weighted arithmetic mean in the calculation of the comparison criterion, so we will have the weighted comparison criterion which is denoted $(W\Delta)$ for the general (total) model of the following equation:

$$W\Delta = W_1 * \Delta_1 + W_2 * \Delta_2 \qquad \dots (14)$$

Whereas
$$W_i$$
 equals:
 $W_i = \frac{n_i}{n_1 + n_2}$; $i = 1,2$

Equation (14) was used in the comparison criteria AIC, MDL, FPE and MSE.

The Practical Part:

Introduction:

Air pollution plays an important role in the air quality management system. The critical periods of the expectations of air pollution are determined by using the relevant weather data in the area with the expectations in the current period. The soils and dust are regarded the main elements in dusty fog in cities, and are among the most complicated pollutants which are difficult to control. This has an important influence on health. In addition, the desert character of the areas surrounding Baghdad caused an increase in the pollution cases of suspended solids, and the increase in storms raised the concentrations of dust in these areas. Further, the low relative humidity and the general average of rain increase the atmospheric lifetime of air pollutants. Consequently, the environmental effects influence human health and cause damage to plants and materials.(20, 21)

Climate data was taken and analyzed after determining the threshold which separates one period and another by using specific calculations, and the ARX model was built using threshold, for monthly data represented by two climate elements, i.e., soils, dust and wind speed, relating to Baghdad Weather Station for the period from January 1956 until December 2012.

Data Preparation

Data of Baghdad Weather Station, include soils and dust as output series y(t) and wind speed variable as input series u(t) for the period from January 1956 until December 2012. The data was taken from the Weather Forecast Service of the Ministry of Transportation in Iraq. The data was analyzed by using the MATLAB V.2013a program.

The two Figs.2 and 3 show the nature of the spread of these two series. It is noteworthy that the cross correlation between the two series equals (0.682). This is evidence that there is a correlation between the two variables of this paper.

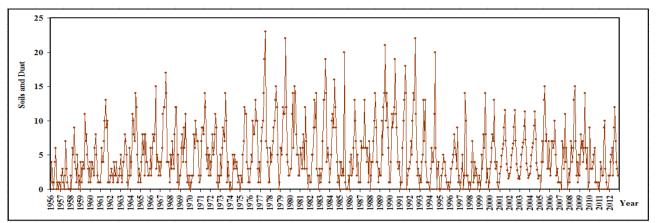


Figure 2. Output series (soils and dust of Baghdad Station in Iraq)

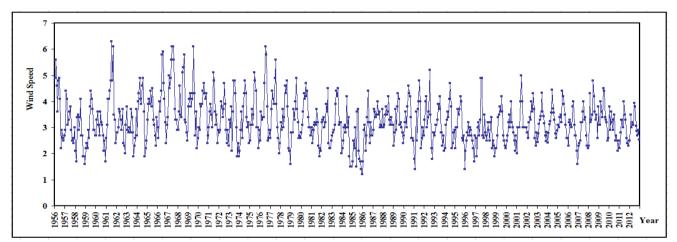


Figure 3. Input series (wind speed of Baghdad Station in Iraq)

While Fig.4 shows the cross correlation between the two series. There is a system in the

relationship between the two variables as is clear from the change in autocorrelation between them.

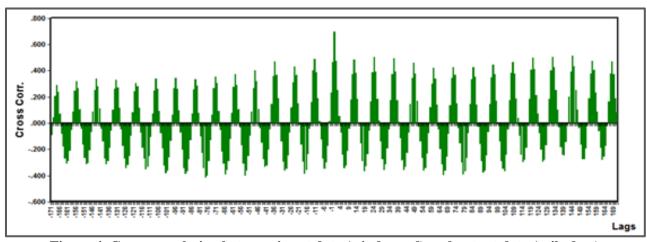


Figure 4. Cross correlation between input data (wind speed) and output data (soils dust)

The First Method: Determining the Best Cut off Point for Data Modeling by Partial Sizes of Samples

To determine the best ARX model against the best cut point of the study data, the check has been started as of the observation (51); March 1960, [depend on create preliminary ARX model from the initial observations, which are 50 observations, symbolization as n_1 and representing from (December 1956 to February 1960) within determining the model's order, post-taking the combinations (from 1 to 5) of the three different parameters (n_a, n_b, n_k)].

Then, searching for the best separation point (threshold point) can be made by dividing the data into two models depending on this point by using some comparison criteria.

This is done through two proposed approaches:

The First Proposed Approach: The Method of Prediction Forward (The Forecasting Approach)

This method depends on the future predictions of the identified models through the following algorithm:

1. Determining the first preliminary separation point, represented by observation 51corresponding to March 1960. Accordingly, the whole sample will be divided into two groups: first group represented by 50 observations, which starts from January 1956 to February 1960, the second group, represented by observations that start with observation 51 until

observation 672, i.e. from March 1969 until December 2012.

- 2. The best model order of first and second groups are determined by comparing the first 12 observations i.e. the year 1956 forecasted from the first group with the actual observations that from 51 to 62 i.e. from March 1960 to February 1961, and also comparing the first 12 values forecasted from the second group with the actual observations that from 673 to 684 i.e. from September 2003 to December 2012. Then, calculate the comparison of the total model.
- 3. The cut point is changed by one observation, namely observation 52 i.e. April 1960 and the items mentioned above are calculated again until the last cut point in observation 622 is reached, namely October 2007, where the sample will divide into two groups, the first from observation 1 to 621 observation i.e. from January 1956 to September 2007, the second group represented by observations from 622 to 672 i.e. from October 2007 to December 2011, and compared with the first 12 values for the two samples to calculate the comparison criteria which are AIC, MDL, Loss Function, and MSE criterion.
- 4. Determining the best cut point and the best model through the smallest measure of comparison to the criteria mentioned in paragraph (3).

The results are as follows: At the AIC criterion:

Table 1. Results of AIC criterion of the first and second groups, and many orders of ARX model, for different sample sizes (The Forecasting Approach)

AIC ₁	na ₁	nb ₁	nk ₁	$\overline{\mathbf{n_1}}$	AIC ₂	na ₂	nb ₂	nk ₂	n ₂	WAIC
1.7655	1	1	2	50	-5.4904	1	4	5	622	-4.9506
-2.3106	2	1	1	52	-5.4909	1	4	5	620	-5.2448
-2.6276	1	3	2	60	-5.7005	1	4	5	612	-5.4261
-4.3528	3	3	3	62	-5.7041	1	4	5	610	-5.5795
-4.4486	1	5	1	71	-5.7408	2	5	4	601	-5.6043
-8.7525	1	4	5	80	-5.9440	2	5	4	592	-6.2783
-6.9916	5	3	1	131	-6.1800	1	5	4	541	-6.3382
-6.5936	5	5	2	496	-5.6604	4	5	5	176	-6.3492
-8.4339	1	3	1	539	-9.1041	5	1	2	133	-8.5666

At the MDL criterion:

Table 2. Results of MDL criterion of the first and second groups, and many orders of ARX model, for different sample sizes (The Forecasting Approach)

	unitation sumple sizes (The 1 of ceasing rippi out in)												
MDL_1	na_1	nb_1	nk_1	\mathbf{n}_1	$\mathbf{MDL_2}$	na_2	nb_2	nk_2	\mathbf{n}_2	WMDL			
1.8394	1	1	2	50	-5.4556	1	4	5	622	-4.9128			
-2.2075	2	1	1	52	-5.4560	1	4	5	620	-5.2046			
-2.5037	1	3	2	60	-5.6652	1	4	5	612	-5.3829			
-4.1848	3	3	3	62	-5.6687	1	4	5	610	-5.5318			
-4.2886	1	5	1	71	-5.6914	2	5	4	601	-5.5431			
-8.6210	1	4	5	80	-5.8940	2	5	4	592	-6.2186			
-6.8403	5	3	1	131	-6.1338	1	5	4	541	-6.2716			
-8.4026	1	3	1	539	-8.9863	5	1	2	133	-8.5181			

At the Loss Function:

		u.	mer em	Sample	Sizes (The Tol	ccasum	ig Appr	vacii)		
Loss Fun. ₁	na ₁	nb ₁	nk ₁	$\overline{n_1}$	Loss Fun. ₂	na ₂	nb ₂	nk ₂	n ₂	WLoss .Fun.
1.6481	1	2	1	50	-5.5066	1	4	5	622	-4.9743
-2.4442	2	1	1	52	-5.5072	1	4	5	620	-5.2702
-2.7760	1	3	2	60	-5.7170	1	4	5	612	-5.4544
-4.5605	3	3	3	62	-5.7207	1	4	5	610	-5.6136
-4.6281	1	5	1	71	-5.7642	2	5	4	601	-5.6442
-8.8860	1	4	5	80	-5.9678	2	5	4	592	-6.3152
-7.1158	5	3	1	131	-6.2023	1	5	4	541	-6.3804
-6.6340	5	5	2	496	-5.7634	4	5	5	176	-6.4060
-8 4489	1	3	1	539	-9 1972	5	1	2.	133	-8 5970

Table 3. Results of Loss Function of the first and second groups, and many orders of ARX model, for different sample sizes (The Forecasting Approach)

The results of the weighted comparison criteria in Tables 1, 2, 3 show that the minimum value is (-8.5666, -8.5181, -8.5970) respectively. This corresponds to:

In the first group of observations n_1 , which corresponds to the minimum value of those criteria, is when the sample size equals (n_1 =539), namely the period from December 1956 to November 2000, at parameters (na=1, nb=3, nk=1) for all comparison criteria used which equal (AIC= -8.4339, MDL= -8.4026, Loss Fun.= -8.4489). The estimation of ARX model equals:

$$A(q) = 1 - 0.5602q^{-1}$$
Input $B(q) = -0.0556q^{-1} + 0.0233q^{-2} + 0.0805q^{-3}$

And MSE = 11.06

In the second group, which corresponds to the minimum value in the weighted comparison criteria, the sample size is equal (n_2 =133), namely the period from December 2000 to December 2011, at parameters (n_2 =5, n_2 =1, n_3 =2) for all comparison criteria used which equal (AIC = -9.1041, MDL = -8.9863, Loss Fun.= -9.1972). The estimation of ARX model equals:

$$A(q) = 1 - 0.6019 q^{-1} + 0.0302 q^{-2} - 0.2183 q^{-3} + 0.1574 q^{-4} + 0.0704 q^{-5}$$

Input
$$B(q) = 0.0485q^{-2}$$

And MSE = 7.745

Therefore, the weighted mean square error (WMSE) of the two groups was calculated and equals (10.404).

The Second Proposed Approach: The Method of Prediction from Within the Time Series (The Prediction Approach)

This approach resembles the first approach (the algorithm mentioned in the first approach) with the difference that the comparison is made by predicting from inside the time series of all the first sample whose size is 50 observations, namely January 1956 to February 1960, the second sample is represented by (n-50) which is equal to 634 observations from March 1960 to December 2012. Then, the cut point is changed on the basis that it moves the following observation, observation 52 i.e. April 1960, and so on until the last best cut point (threshold) of ARX model is reached.

The results are as follows:

At the AIC criterion:

Table 4. Results of AIC criterion of the first and second groups, and many orders of ARX model, for different sample sizes (The Prediction Approach)

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AIC ₁	na_1	$\mathbf{nb_1}$	nk_1	$\mathbf{n_1}$	AIC_2	na_2	nb_2	nk_2	\mathbf{n}_2	WAIC
1.6959	3	1	5	50	2.1970	5	5	5	634	2.1603
1.6852	3	1	4	51	2.1986	5	5	5	633	2.1603
1.6733	3	1	5	52	2.1967	5	5	5	632	2.1569
1.6228	4	4	5	57	2.2041	5	5	5	627	2.1556
1.5965	4	4	5	58	2.2058	5	5	5	626	2.1541
1.5716	4	4	5	59	2.2029	5	5	5	625	2.1484
1.5466	4	4	5	60	2.2011	5	5	5	624	2.1437
1.5223	4	4	5	61	2.2027	5	5	5	623	2.1420
1.6527	4	4	5	75	2.2022	5	5	5	609	2.1420
1.6806	3	1	5	86	2.2068	5	5	5	598	2.1407
1.6702	3	1	5	88	2.2092	5	5	5	596	2.1398
1.6572	3	1	5	89	2.2057	5	5	5	595	2.1344
1.6472	4	4	5	91	2.2081	5	5	5	593	2.1335
1.6339	4	4	5	92	2.2079	5	5	5	592	2.1307
1.6391	4	4	5	93	2.2048	5	5	5	591	2.1279
1.6354	4	4	5	94	2.2050	5	5	5	590	2.1268
1.6229	4	4	5	95	2.2067	5	5	5	589	2.1256
1.6171	4	4	5	98	2.2103	5	5	5	586	2.1253

From Table 4 it is obvious that the minimum value of the weighted AIC criterion of data equals (WAIC=2.1253). This corresponds to: In the first group the sample size is equal (n=98), whereas (AIC=1.6171), when the number of parameters is (na=4, nb=4, nk=5). And the estimation of the ARX model is:

$$A(q) = 1 - 0.3959q^{-1} - 0.1615q^{-2} + \\ 0.2500q^{-3} - 0.1734q^{-4}$$
 Input
$$B(q) = 0.0258q^{-5} + 0.0102q^{-6} + \\ 0.0747q^{-7} - 0.0724q^{-8}$$

And MSE = 4.997.

In the second group of observations, which corresponds to the least WAIC, the sample size is equal $(N-n_1=586)$ and the criterion value is

(AIC=2.2103), when the number of parameters is (na=5, nb=5, nk=5). And the estimation of the ARX model is:

$$A(q) = 1 - 0.3030q^{-1} - 0.0435q^{-2} - 0.0845q^{-3} - 0.0534q^{-4} - 0.0716q^{-5}$$
Input
$$B(q) = 0.0316q^{-5} + 0.1064q^{-6} + 0.0012q^{-7} - 0.0299q^{-8} - 0.0577q^{-9}$$

And MSE = 8.894.

The WMSE of the two groups equals (8.336).

At the MDL criterion:

Table 5 shows that the minimum value of the weighted MDL criterion equals (WMDL=2.1928). This corresponds to:

Table 5. Results of MDL criterion of the first and second groups, and many orders of ARX model, for different sample sizes (The Prediction Approach)

$\overline{\mathbf{MDL_1}}$	na ₁	nb ₁	nk ₁	\mathbf{n}_1	$\overline{\mathrm{MDL}_2}$	na ₂	nb ₂	nk ₂	\mathbf{n}_2	WMDL
1.8071	1	1	5	50	2.2549	1	4	5	634	2.2222
1.7859	1	1	5	51	2.2566	1	4	5	633	2.2215
1.7640	1	1	5	52	2.2558	1	4	5	632	2.2184
1.7142	1	1	5	58	2.2648	1	4	5	626	2.2182
1.6926	1	1	5	59	2.2626	1	4	5	625	2.2134
1.6712	1	1	5	60	2.2601	1	4	5	624	2.2084
1.6510	1	1	5	61	2.2618	1	4	5	623	2.2073
1.7579	1	1	5	84	2.2699	1	4	5	600	2.2070
1.7563	1	1	5	86	2.2689	1	4	5	598	2.2045
1.7478	1	1	5	87	2.2707	1	4	5	597	2.2042
1.7373	1	1	5	88	2.2724	1	4	5	596	2.2035
1.7254	1	1	5	89	2.2694	1	4	5	595	2.1986
1.7145	1	1	5	92	2.2724	1	4	5	592	2.1974
1.7263	1	1	5	93	2.2672	1	4	5	591	2.1936
1.7227	1	1	5	95	2.2687	1	4	5	589	2.1928

In the first group, the sample size is equal (n=95), and (MDL=1.7227), when the number of parameters is (na=1, nb=1, nk=5). And the estimation of the ARX model is:

$$A(q) = 1 - 0.3884q^{-1}$$

Input $B(q) = 0.0460q^{-5}$

And MSE = 5.049

As for the results of the second group, it will be at the sample size $(N-n_1=589)$ which corresponds to (MDL=2.2687) when the number of parameters

(na=1,nb=4, nk=5). And the estimation of the ARX model is:

$$A(q) = 1 - 0.3797q^{-1}$$
Input
$$B(q) = 0.0195q^{-5} + 0.1170q^{-6} - 9.387E - 05q^{-7} - 0.0628q^{-8}$$

And MSE = 9.182

The WMSE of the two groups was calculated and equals (8.608).

At the Loss Function:

Table 6. Results of Loss Function criterion of the first and second groups, and many orders of ARX
model, for different sample sizes (The Prediction Approach)

Loss Fun. ₁	na ₁	nb ₁	nk ₁	n ₁	Loss Fun. ₂	na ₂	nb ₂	nk ₂	n ₂	WLoss. Fun.
1.3588	5	5	5	50	2.1654	5	5	5	634	2.1065
1.3550	5	5	4	51	2.1670	5	5	5	633	2.1064
1.3629	5	5	4	52	2.1650	5	5	5	632	2.1040
1.3518	5	5	4	53	2.1666	5	5	5	631	2.1035
1.3353	5	5	4	54	2.1681	5	5	5	630	2.1024
1.3250	5	5	4	56	2.1712	5	5	5	628	2.1019
1.3116	5	5	4	57	2.1722	5	5	5	627	2.1005
1.2906	5	5	4	58	2.1738	5	5	5	626	2.0989
1.2711	5	5	4	59	2.1709	5	5	5	625	2.0933
1.2509	5	5	4	60	2.1691	5	5	5	624	2.0885
1.2313	5	5	4	61	2.1706	5	5	5	623	2.0868
1.4129	5	5	5	75	2.1694	5	5	5	609	2.0864
1.4794	5	5	5	88	2.1756	5	5	5	596	2.0860
1.4673	5	5	5	89	2.1721	5	5	5	595	2.0804
1.4547	5	5	5	91	2.1744	5	5	5	593	2.0786
1.4435	5	5	5	92	2.1741	5	5	5	592	2.0759
1.4492	5	5	5	93	2.1709	5	5	5	591	2.0728
1.4465	5	5	5	94	2.1711	5	5	5	590	2.0716
1.4354	5	5	5	95	2.1727	5	5	5	589	2.0703
1.4378	5	5	4	97	2.1745	5	5	5	587	2.0700
1.4306	5	5	4	98	2.1762	5	5	5	586	2.0694
2.2203	5	5	5	489	1.6884	5	5	5	195	2.0687

Table 6 shows that the minimum value of the weighted loss function (WLoss Fun.) is equal (2.0687) which corresponds to the following:

In the first group of data has a sample size of (n=489) and (Loss. Fun.= 2.2203), when the number of parameters is (na=5, nb=5, nk= 5). And the estimation of ARX model is equal to:

$$A(q) = 1 - 0.3376q^{-1} - 0.0951q^{-2} - 0.0237q^{-3} - 0.0930q^{-4} - 0.0557q^{-5}$$
 Input
$$B(q) = 0.0368q^{-5} + 0.0842q^{-6} + 0.0269q^{-7} - 0.0637q^{-8} - 0.0392q^{-9}$$

And MSE = 9.364

As for the second group of data, it will be at the sample size $(N-n_1=195)$ and (Loss. Fun.=1.6884), when the number of parameters (na=5, nb=5, nk=5). And the estimation of ARX model is equal to:

$$A(q) = 1 - 0.3117q^{-1} + 0.1579 q^{-2} - 0.2047q^{-3} + 0.0347q^{-4} - 0.1220q^{-5}$$

Input
$$B(q) = 0.0252q^{-5} + 0.1169q^{-6} - 0.0647q^{-7} + 0.0858q^{-8} - 0.1050q^{-9}$$

And MSE = 5.995

And the value of WMSE for the results of two groups is equal to (8.404)

The second method: Determine the Size of the Sample Using the Seasonal B-J Models

From Fig.2 for the soils and dust series of the Baghdad Station in Iraq, and Fig.5 for the autocorrelation (AC) and the partial autocorrelation (PAC) respectively by using SPSS V.20 program, it is clear that there is no stationary due to the sharp curves in the series data.

In the Dickey-Feller test, it shows stationary in the time series, although the series is found to be non-stationary when looking at the AC and PAC. Indicating the ratio of outliers values in the series so that they do not affect the total size of the sample in terms of stationary. Where the p-value is less than 0.05 for all test models shown in the Table 7, this confirms that the series is stationary.

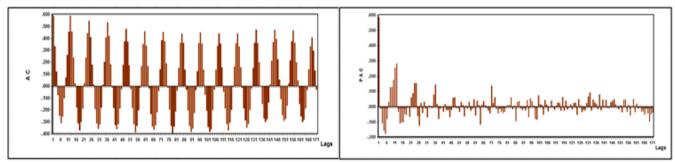


Figure 5. Coefficients of AC and PAC for the output series (soils and dust) of Baghdad Station

Table 7. Dickey-Feller test for soils and dust series of Baghdad Station

Test	P-Val	ue
	Output	Input
Constant	8.735E-006	0.0003
Constant and trend	8.1296E-005	0.0008

The most important tests used in seasonal pattern detection are Kruskal-Wallis and Jonckheere-Terpstra tests. The hypothesis of the test is as follows:

H₀: There is no seasonal in the data

H₁: There is seasonal in the data

The results are as follows:

Table 8. Kruskal-Wallis and Jonckheere-Terpstra test of soils and dust data at Baghdad Station

Test St	atistics ^{a,b}	Jonckheere-Terps	tra Test ^a		Hypothesis Test Summary					
	soils_dust		soils_dust	Ш	Null Hypothesis	Test	Sig.	Decision		
Chi-Square df	318.654 11	Number of Levels in MONTH, period 12	12		The distribution of soils_dust is th 1 same across categories of MONT period 12.	e Independent- Samples H'Kruskal-Wallis Test	.000	Reject the null hypothesis.		
Asymp. Sig.	.000	N	684	н		Independent-				
a. Kruskal Wallis To b. Grouping Variab	est le: MONTH, period 12	Observed J-T Statistic Mean J-T Statistic Std. Deviation of J-T	96187.000 107217.000	1	The distribution of soils_dust is th 2 same across categories of MONT period 12.			Reject the null hypothesis.		
		Statistic Statistic	2960.791		Asymptotic significances are display	ed. The significan	ce level	is .05.		
		Std. J-T Statistic	-3.725	ı						
		Asymp. Sig. (2-tailed)	.000	J						
		a. Grouping Variable: MO	NTH, period 12							

The results of the Table 8 show that the null hypothesis is rejected at a significant level of 0.05, i.e., the seasonal data containment that occurs every 12 months.

Thus, a number of seasonal B-J models (i.e. seasonal autoregressive integrated moving average SARIMA model) have been applied to the data using MATLAB V.2013a program, through two approaches:

The First Approach: (The Forecasting Approach)

In this approach, a 7560 seasonal models are tested. The program has been implemented continuously (i.e. is the implementation period of the program and not the period of the building). The best order were selected according to the minimum value of the comparison criteria using the forecasting principle of 12 observations forward, i.e. adopt a sample size of 672 out of 684 observations. This technique can be called a (Sample Test). The results are shown in Table 9.

Table 9. The significant models of seasonal B-J with the corresponding comparison criteria by testing a sample of 672 observations by (Forecasting Approach) (Sample Test)

a sample of 0/2 observations by (Forecasting Approach) (Sample Test)											
n ₁	P	q	P	Q	s	D	D	AIC ₁	BIC ₁	FPE_1	MDL_1
12	0	2	3	0	12	0	0	3648.3771	3656.6205	21.0409	19.6781
12	0	2	2	1	12	0	0	3648.5515	3656.7949	21.8783	20.4612
12	0	3	3	0	12	0	0	3649.3823	3657.1107	23.7083	22.4155
12	0	1	1	0	12	0	0	3660.8691	3667.6578	22.7956	20.5111
12	0	0	1	1	12	0	0	3660.8691	3667.6578	22.7956	20.5111
12	0	2	1	0	12	0	0	3662.6388	3669.9124	23.5745	21.5111
12	1	0	0	2	12	0	0	3662.6388	3669.9124	23.5745	21.5111
12	0	0	2	1	12	0	0	3662.6732	3669.9468	23.5859	21.5215
12	0	1	2	0	12	0	0	3662.6732	3669.9468	23.5859	21.5215
12	0	0	4	0	12	0	0	3663.0645	3670.8230	24.7319	22.8587
12	1	0	3	0	12	0	0	3663.0645	3670.8230	24.7319	22.8587
12	1	0	0	3	12	0	0	3663.7057	3671.4642	24.6390	22.7728
12	0	1	1	2	12	0	0	3663.7057	3671.4642	24.6390	22.7728
12	0	1	3	0	12	0	0	3663.7312	3671.4897	24.5763	22.7149
12	1	3	1	0	12	0	0	3663.9369	3672.1803	25.4360	23.7885
12	0	0	3	0	12	0	0	3664.1872	3671.4608	24.0466	21.9419
12	0	2	2	0	12	0	0	3664.4509	3672.2094	24.4504	22.5985
12	1	0	1	2	12	0	0	3664.5183	3672.2768	24.6324	22.7667
12	1	1	1	1	12	0	0	3664.6565	3672.4150	24.5678	22.7070
12	1	1	2	0	12	0	0	3664.6646	3672.4231	24.4563	22.6039
12	0	1	2	1	12	0	0	3664.8449	3672.6034	24.5588	22.6987
12	0	1	1	1	12	0	0	3665.1048	3672.3784	23.9804	21.8815
12	0	0	0	4	12	0	0	3666.7482	3674.5067	25.0061	23.1121
12	0	0	2	0	12	0	0	3667.8734	3674.6621	23.2323	20.9040
12	1	0	2	0	12	0	0	3668.8135	3676.0871	24.1036	21.9939
12	0	3	0	1	12	0	0	3672.6654	3680.4239	25.1302	23.2268
12	0	0	1	0	12	0	0	3675.1659	3681.4697	22.4753	19.9148
12	1	0	1	1	12	0	0	3679.1671	3686.4407	24.2733	22.1488
12	1	0	1	0	12	0	0	3680.9307	3687.7194	23.4274	21.0796

In the Table 9, it is found that the minimum values for the comparison criteria in the SARIMA $(0,0,2)(3,0,0)_{12}$ model in which the values of: AIC= 3648.3771, BIC= 3656.6205, FPE= 21.0409, MDL= 19.6781 .

Thus, this seasonal model was selected for (Sample Testing), where the value of MSE= 12.1850

The estimation of the parameters of this seasonal model is shown in the Table 10:

When testing the accuracy of the model, the following Fig.6 illustrates the randomized model and AC and PAC coefficients of the residuals are located within the confidence interval and equal to $\{\pm 0.075\}$.

Table 10. Estimating the parameters of the SARIMA(0,0,2)(3,0,0)₁₂ model in (Forecasting Approach/Sample Test)

Parameters	Estimate	S.E	t-statistic	Sig.
MA(1)	-0.292	0.039	-7.587	0.000
MA(2)	-0.176	0.039	-4.556	0.000
SAR(1)	0.328	0.038	8.726	0.000
SAR(2)	0.273	0.039	7.055	0.000
SAR(3)	0.265	0.038	6.928	0.000

The value of the statistic Q $_{\text{Ljung-Box}}$ = 45.444 and compared with the value of χ^2 tabular and degree of freedom (45) and significant level (0.05) is equal to (61.3850), it is clear that the model is appropriate to represent the data.

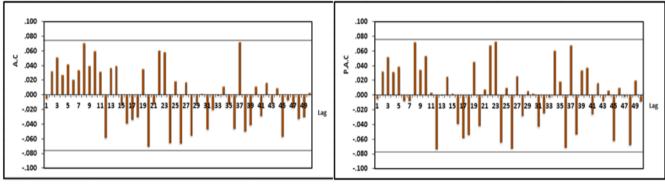


Figure 6. AC and PAC coefficients for the residuals of the SARIMA(0,0,2)(3,0,0)₁₂ model

The Second Approach: (The Prediction Approach)

In this method, 7560 seasonal models were tested through test 684 observations as a sample, these observations are predicted within the series and all the criteria are shown to be consistent with a seasonal model (s = 12).

The combinations (0 to 5) were taken for each of the four parameters model (p, q, P, Q), and the results are shown in Table 11 and this technique can be called a (Sample Estimation).

Table 11. The significant models of seasonal B-J with the corresponding comparison criteria by testing a sample of 684 observations in (Prediction Approach) (Sample Estimation)

	•	Juiipi	U UI U	0.000	or tau	OIID II	1 (1 1 0	diction rippr	oach) (Bampi	2 Estimation	· /
n	P	Q	P	Q	S	d	D	AIC	BIC	FPE	MDL
684	0	4	5	0	12	0	0	3603.7173	3644.4689	22.4349	12.0228
684	0	5	4	0	12	0	0	3615.8091	3656.5607	22.8351	12.2372
684	0	3	3	0	12	0	0	3645.9703	3673.1380	23.9705	12.5614
684	0	4	4	0	12	0	0	3641.1931	3677.4168	23.7335	12.6251
684	0	5	0	5	12	0	0	3635.6034	3680.8830	23.4707	12.6701
684	0	3	5	0	12	0	0	3644.7906	3681.0143	23.8587	12.6917
684	0	4	0	5	12	0	0	3642.5278	3683.2794	23.7447	12.7247
684	0	5	0	4	12	0	0	3642.5278	3683.2794	23.7447	12.7247
684	3	0	0	2	12	0	0	3661.2492	3683.8890	24.5482	12.7665
684	0	2	3	0	12	0	0	3661.2492	3683.8890	24.5482	12.7665
684	0	4	0	4	12	0	0	3653.0761	3689.2997	24.1494	12.8464
684	2	4	1	0	12	0	0	3661.2786	3692.9743	24.4769	12.9238
684	3	0	0	4	12	0	0	3661.2786	3692.9743	24.4769	12.9238
684	0	4	3	0	12	0	0	3661.2786	3692.9743	24.4769	12.9238
684	0	3	2	0	12	0	0	3672.1182	3694.7580	24.9414	12.9710
684	2	2	5	0	12	0	0	3658.4483	3699.2000	24.3039	13.0243
684	0	1	1	0	12	0	0	3692.8467	3701.9027	25.8224	13.1193
684	2	0	0	5	12	0	0	3672.8631	3704.5588	24.8950	13.1446
684	3	0	0	5	12	0	0	3668.8797	3705.1034	24.7139	13.1466

684	0	5	3	0	12	0	0	3668.8797	3705.1034	24.7139	13.1466
684	5	3	2	0	12	0	0	3663.1940	3708.4736	24.4368	13.1916
684	0	3	4	0	12	0	0	3676.0889	3707.7846	25.0127	13.2067
684	1	0	0	2	12	0	0	3694.5792	3708.1630	25.8499	13.2369
684	0	2	1	0	12	0	0	3694.5792	3708.1630	25.8499	13.2369
684	0	3	0	5	12	0	0	3673.7925	3710.0161	24.8920	13.2414
684	0	1	1	1	12	0	0	3697.2159	3710.7998	25.9497	13.2881
684	2	2	4	0	12	0	0	3677.2593	3713.4830	25.0185	13.3087
684	0	3	0	4	12	0	0	3681.6078	3713.3035	25.2153	13.3137
684	0	2	0	5	12	0	ő	3681.6078	3713.3035	25.2153	13.3137
684	0	3	1	0	12	0	0	3695.6121	3713.7239	25.8509	13.3409
684	0	1	3	0	12	0	0	3695.6298	3713.7237	25.8516	13.3412
684	1	1	1	1	12	0	0	3696.6295	3714.7413	25.8894	13.3607
684	1	0	0	4	12	0	0	3696.7342	3719.3740	25.8553	13.4463
	0			-							
684		4	1	0	12	0	0	3696.7342	3719.3740	25.8553	13.4463
684	3	0	1	1	12	0	0	3696.9404	3719.5802	25.8631	13.4503
684	2	0	3	0	12	0	0	3696.9457	3719.5855	25.8633	13.4504
684	1	0	4	0	12	0	0	3696.9457	3719.5855	25.8633	13.4504
684	0	1	0	2	12	0	0	3706.0875	3719.6713	26.2885	13.4615
684	1	1	0	3	12	0	0	3697.6516	3720.2913	25.8900	13.4643
684	1	2	0	4	12	0	0	3690.1108	3721.8065	25.5307	13.4802
684	0	3	0	2	12	0	0	3699.7753	3722.4150	25.9705	13.5062
684	0	4	0	1	12	0	0	3699.7753	3722.4150	25.9705	13.5062
684	0	1	0	4	12	0	0	3699.7753	3722.4150	25.9705	13.5062
684	0	2	0	3	12	0	0	3699.7753	3722.4150	25.9705	13.5062
684	1	0	1	0	12	0	0	3713.5532	3722.6091	26.6160	13.5225
684	0	1	0	3	12	0	0	3704.9860	3723.0979	26.2077	13.5250
684	0	3	0	1	12	0	0	3704.9860	3723.0979	26.2077	13.5250
684	2	0	0	4	12	0	0	3698.4942	3725.6620	25.8837	13.5640
684	0	4	2	0	12	0	0	3698.4942	3725.6620	25.8837	13.5640
684	1	0	0	5	12	0	0	3698.5489	3725.7167	25.8858	13.5651
684	0	5	1	0	12	0	0	3698.5489	3725.7167	25.8858	13.5651
684	1	1	0	4	12	0	0	3698.5489	3725.7167	25.8858	13.5651
684	0	4	1	1	12	0	0	3698.5489	3725.7167	25.8858	13.5651
684	0	2	0	4	12	0	0	3698.6470	3725.8148	25.8895	13.5670
684	0	4	0	2	12	0	0	3698.6470	3725.8148	25.8895	13.5670
684	2	0	2	0	12	0	0	3707.1556	3725.2675	26.2909	13.5679
684	0	2	4	0	12	0	0	3698.7945	3725.9622	25.8951	13.5699
							_	3711.6441	3725.2280		
684	1	0	1	1	12	0	0			26.5029	13.5713
684	1	1	1	0	12	0	0	3711.6441	3725.2280	26.5029	13.5713
684	2	0	4	0	12	0	0	3698.8995	3726.0672	25.8991	13.5720
684	4	0	1	1	12	0	0	3698.9378	3726.1055	25.9005	13.5728
684	1	1	4	0	12	0	0	3698.9378	3726.1055	25.9005	13.5728
684	1	0	5	0	12	0	0	3698.9440	3726.1118	25.9008	13.5729
684	3	0	3	0	12	0	0	3698.9440	3726.1118	25.9008	13.5729
684	0	1	5	0	12	0	0	3698.9559	3726.1237	25.9012	13.5731
684	0	1	0	5	12	0	0	3701.7103	3728.8780	26.0057	13.6279
684	0	3	0	3	12	0	0	3701.7103	3728.8780	26.0057	13.6279
684	0	5	0	1	12	0	0	3701.7103	3728.8780	26.0057	13.6279
684	0	1	0	1	12	0	0	3720.3986	3729.4545	26.8837	13.6585
684	1	1	0	5	12	0	0	3699.7477	3731.4434	25.8930	13.6715
684	0	2	0	2	12	0	0	3712.6889	3730.8008	26.5045	13.6781
684	4	0	3	0	12	0	0	3700.6885	3732.3842	25.9286	13.6903
684	2	0	5	0	12	0	0	3700.6885	3732.3842	25.9286	13.6903
684	3	0	4	0	12	0	0	3700.6885	3732.3842	25.9286	13.6903
684	0	2	5	0	12	0	0	3700.7907	3732.4864	25.9325	13.6924
684	1	1	5	0	12	0	0	3700.9358	3732.6315	25.9380	13.6953
684	4	0	4	0	12	0	0	3701.2822	3737.5058	25.9128	13.7844
684	3	0	5	0	12	0	0	3701.2822	3737.5058	25.9128	13.7844
684	4	0	5	0	12	0	0	3701.5936	3742.3452	25.8863	13.8723
	т	<u> </u>		<u> </u>	. 4		J	3101.3730	3114,3734	25.0005	13.0123

The Table 11 shows that the minimum values of the comparison criteria in the SARIMA $(0,0,4)(5,0,0)_{12}$ model, where: AIC= 3603.7173, BIC= 3644.4689, FPE= 22.4349, MDL= 12.0228

Thus, this seasonal model was selected for (Sample Estimation), where the value of MSE=11.0718 The estimation of the parameters of this seasonal model is given in the Table 12:

Table 12. Estimating the parameters of the SARIMA $(0,0,4)(5,0,0)_{12}$ model for (Prediction Approach) (Sample Estimation)

Parameters	Estimate	S. E	t-statistic	Sig.
MA(1)	-0.291	0.038	-7.596	0.000
MA(2)	-0.209	0.039	-5.321	0.000
MA(3)	-0.183	0.039	-4.666	0.000
MA(4)	-0.172	0.038	-4.505	0.000
SAR(1)	0.248	0.038	6.459	0.000
SAR(2)	0.204	0.040	5.171	0.000
SAR(3)	0.180	0.040	4.494	0.000
SAR(4)	0.103	0.040	2.564	0.009
SAR(5)	.1420	.0400	3.593	.0000

When testing the accuracy of the model, the Fig.7 illustrates the randomized model and AC, PAC coefficients of the residuals are within the confidence interval and its: $\{\pm 0.075\}$

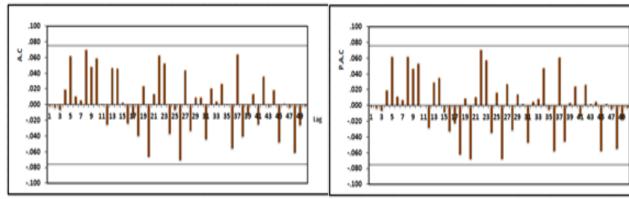


Figure 7. AC and PAC coefficients for the residuals of the SARIMA(0,0,4)(5,0,0)₁₂ model

By comparing the value Q $_{\text{Ljung-Box}} = 21.107$ with χ^2 tabular at a d.f = 41 and a significant level of 0.05 it's equal to (56.9350), it is clear that the model is appropriate to represent the data.

Conclusions:

- 1- All the comparison criteria used in the forecasting method agree on the size of a sample at a certain threshold and agreed on the same number of parameters used to estimate the model at each threshold. Thus, this method can be considered more stable in determining the threshold point and the order of the ARX model.
- 2- There is a slight variation in the results of the comparison criteria (AIC and MDL) used in the predictive method within the series in terms of sample size at a certain threshold, as opposed to the results of the loss function in which a large difference in sample size appears at a given threshold. The number of parameters shows the specific values of the orders differed according to the criteria used.
- 3- The last threshold changes the function when the first approach (Forecasting) is at the threshold

- size of the sample 539, then the ARX model is stabilized and unchanged until the last observation.
- 4- The results show that the best cut-off model is determined by the second approach (Prediction) at the AIC criterion where WMSE = 8.336 followed by the use of the loss function in which WMSE = 8.404, while MLD criterion is the last and WMSE = 8.608.
- 5- Through the results of the weighted mean square error, it is clear that the second approach is the best in determining the best threshold point for modeling and equal to 8.336 while its value is equal to 10.404 when using the first approach.
- 6- When using the first approach of sample determination using seasonal B-J models, the SARIMA(0,0,2)(3,0,0)₁₂ model is the best fit for time series data depending on a sample of 672 out of 684 observations (according to the principle of forecasting the 12-observation forward).
- 7- When the second approach is used to determine sample size using seasonal B-J models, the SARIMA(0,0,4)(5,0,0)₁₂ model is

- the best fit for time series data depending on the total sample size.
- 8- When comparing the results of the two approaches of the seasonal B-J models, it is clear that the second approach in which MSE=11.0718 is better than using the first approach in which MSE=12.1850.
- 9- When comparing all the results of the research, it is clear that the first method (i.e., determining the best cut off point for data modeling by partial sizes of samples) is preferable to the second method, which includes sample size determination using seasonal B-J models.
- 10- The results of the research also indicate that the second approach (Prediction) is the best in determining the best threshold point during which the data can be divided into two models through that point. This indicates that the approach of prediction from within the series of threshold ARX model is the best performed compared to the SARIMA models.

Recommendations

- 1- From the results, more than one threshold has been maintained for a long time. Therefore, we recommend studying the case of more than a threshold within one group of time series data.
- 2- By drawing the time series, some values were observed beyond the limits of control. The Dickey-Fuller test found that the series in its general form was stationary. This indicates that the ratio of the number of outliers to the size of the sample has a direct effect on the decision that the series is stationary or not. Therefore, we recommend studying the effect of the number of outliers' values on the size of the sample and its effect on the stationary decision of the time series.
- 3- Using nonlinear models for the threshold ARX model.
- 4- Using a multiple inputs with single output system (MISO) with the threshold problem.
- 5- Using a multiple inputs with multiple outputs system (MIMO) with the threshold problem.
- 6- Using the threshold ARX model on the basis of each section (sample) on both sides of the threshold point is represented by a seasonal model and with determined orders.
- 7- For a large number of sources of air pollution from the environment, domestic and industrial...., so a particular network must be done to monitor the type of pollution in a particular area.

References:

- 1. Chatfield C. The analysis of time series an introduction. 6th ed. Chapman & Hall/CRC. By Taylor& Francis Group.2003.
- 2. Franses PH, Dijk DV. Nonlinear time series models in empirical finance. Cambridge University Press. 2000.
- Granger CWJ, Terasvirta T. Modelling nonlinear economic relationships. Oxford University Press. 1993.
- Chen R, Tsay RS. Nonlinear additive ARX models, Journal of the American Statistical Association. 1993 September; 88(423):955-967.
- 5. Lucheroni C. TARX models for spikes and antispikes in electricity markets. Conference Paper in SSRN Electronic Journal. Source: IEEE Xplore. 2010 July: 1-7. DOI:10.1109/EEM.2010.5558724.
- 6. Frimpong JM, Oteng-Abayie EF. When is inflation harmful? Estimating the threshold effect for Ghana. *American Journal of Economics and Business Administration*. 2010; 2 (3): 232-239.
- Filipovic V, Stojanovic V, Nedic N, Prsic D. TARX model of pneumatic cylinder and identification. Conference Paper: Conference: SAUM, At Nis, Serbia, V.: XI International Conference on Systems, Automatic Control and Measurements. 2012 November; 260-263.
- 8. Tsay RS. Analysis of financial time series. John Wiley & Sons, 2nd edition, New-Jersey. 2001.
- Tong, H. Nonlinear time series: A dynamical system approach. Oxford University Press Inc. New York.1990
- Li Z. Autoregression models for trust management in wireless Ad Hoc networks. Thesis, School of Electrical Engineering & Computer Science, University of Ottawapp, Canada. 2011:1-84.
- 11. Ohlsson H , Ljung L, Boyd S. Segmentation of ARX-models using sum-of-norms regularization. Automatica. 2010; 46(6):1107-1111.June.
- 12. Knotters M, Bierkens MFP. Physical basis of time series models for water table depths. Water Resources. Research 2000 Janury; 36(1):181-188.
- 13. Al-Talib M S. Employing ridge regression technique in prediction of the black box models with application. Iraqi Journal of Statistical Science.2012; 22: 121-135.
- 14. Kapetanios, G. Threshold models for trended time series. Empirical Economics. 2003; 28(4):687-707. Springer. DOI 10.1007/s00181-003-0154-8
- 15. Cs´aji BC, Weyer E. Recursive estimation of ARX systems using binary sensors with adjustable thresholds. Preprints of the 16th IFAC Symposium on System Identification, Brussels, Belgium. 2012July;45(16):1185-1190.
- Weron R, Misiorek A. Short-term electricity price forecasting with time series models: a review and evaluation. Series: The European power Supply Industry-Complex Electricity Markets. 2006:231-254.
- 17. Boardman A, Schlindwein FS, Rocha AP, Leite A. A study on the optimum order of autoregressive models for heart rate variability. Institute of Physics Publishing. 2002; 23: 325-336.

- 18. Akaike H . A bayesian extension of the minimum AIC procedure of autoregressive model fitting. Biometrika.1979 August; 66(2):237-242.
- Makridakis S, Wheelwright SC, Hyndman R. Forecasting methods and application. John Wiley & Sons, Inc. New-York. 1998
- 20. Al-Tahi W KJ. Environmental pollution and green economy. Ministry of finance. Economic circle,
- Economic policy department.1-25. http://www.mof.gov.iq/Lists/ResearchesAndStudies/t louth.pdf
- 21. Kim SE, Kumar A. Accounting seasonal nonstationarity in time series models for short-term ozone level forecast. Stochastic Environmental Research and Risk Assessment. 2005 October; 19(4): 241–248.

طريقة مطَّورة باستخدام التنبؤ في بناء افضل انموذج عتبة لـ ARX مع تطبيق عملي 2 احلام احمد جمعة 1

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لخلاصة :

تعتبر النماذج غير الخطية من الطرق المهمة في تحليل السلاسل الزمنية والتي تتميز بامكانيتها الواسعة في عملية التنبؤ لمختلف المظواهر منها الفيزيائية والهندسية والاقتصادية، من خلال دراسة خصائص الاضطرابات العشوائية فيها للتوصل الى التنبؤات بشكل دقيق. المظواهر منها الفيزيائية والهندسية والاقتصادية، من خلال دراسة خصائص الاضطرابات العشوائية فيها للتوصل الى التنبؤ مقترحين تم توظيفهما لغرض تحديد افضل نقطة قطع (عتبة) هما [التنبؤ الى الامام (Forecasting) والتنبؤ من داخل السلسلة الزمنية (Prediction) من خلال مؤشر نقطة العتبة]. بالاضافة الى استخدام نماذج B-J الموسمية الاعتيادية كطريقة ثانية اعتماداً على مبدأ الاسلوبين المقترحين في تحديد افضل انموذج موسمي. والمقارنة مع النماذج المستحصلة عليها من الاسلوبين المذكورة اعلاه للطريقتين، من خلال مجموعة من المعابير وهي Weighted Comparison ، من خلال مجموعة المعابير وهي MSE ، FPE ، BIC ، Loss Function، MDL ، AIC بالاضافة الى معابير المقارنة الموزون الغبار كمتغير مخرج والخاصة بعداد للفترة من شهر كانون الثاني 1956 ولغاية شهر كانون الاول 2012.

الكلمات المفتاحية: ARX، التنبؤ المستقبلي، التنبؤ، الانحدار الذاتي-المتوسط المتحرك المتكامل، العتبة.